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DRAFT: HOW TO TAKE INTO ACCOUNT MODEL INACCURACY WHEN ESTIMATING THE UNCERTAINTY OF THE RESULT OF DATA PROCESSING

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ABSTRACT

In engineering design, it is important to guarantee that the values of certain quantities such as stress level, noise level, vibration level, etc., stay below a certain threshold in all possible situations, i.e., for all possible combinations of the corresponding internal and external parameters. Usually, the number of possible combinations is so large that it is not possible to physically test the system for all these combinations. Instead, we form a computer model of the system, and test this model. In this testing, we need to take into account that the computer models are usually approximate. In this paper, we show that the existing techniques for taking model uncertainty into account overestimate the uncertainty of the results. We also show how we can get more accurate estimates.

INTRODUCTION

Bounds on unwanted processes: an important part of engineering specifications. An engineering system is designed to perform certain tasks. In the process of performing these tasks, the system also generates some undesirable side effects: it can generate noise, vibration, heat, stress, etc.

We cannot completely eliminate these undesired effects, but specifications for an engineering system usually require that the size q of each of these effects does not exceed a certain pre-

defined threshold (bound) t . It is therefore important to check that this specification is always satisfied, i.e., that $q \leq t$ in all possible situations.

How can we check that specifications are satisfied for all possible situations: simulations are needed. To fully describe each situation, we need to know the values of all the parameters p_1, \dots, p_n that characterize this situation.

These may be external parameters such as wind speed, load, etc., for a bridge. This may be internal parameters such as the exact value of the Young module for a material used in the design.

For each of these parameters, we know the interval of possible values $[p_i, \bar{p}_i]$. For many parameters p_i , this interval is described by setting a nominal value \tilde{p}_i and the bound Δ_i on possible deviations from this nominal value. In such a setting, the interval of possible values has the form

$$[p_i, \bar{p}_i] = [\tilde{p}_i - \Delta_i, \tilde{p}_i + \Delta_i]. \quad (1)$$

In other cases, the bounds p_i and \bar{p}_i are given directly. However, we can always describe the resulting interval in the form (1) if we take the midpoint of this interval as \tilde{p}_i and its half-width as Δ_i :

$$\tilde{p}_i \stackrel{\text{def}}{=} \frac{p_i + \bar{p}_i}{2}; \quad \Delta_i \stackrel{\text{def}}{=} \frac{\bar{p}_i - p_i}{2}. \quad (2)$$

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Thus, without losing generality, we can always assume that the set of possible values of each parameter p_i is given by the expression (1).

We would like to make sure that the quantity q satisfies the desired inequality $q \leq t$ for *all* possible combinations of values $p_i \in [\underline{p}_i, \bar{p}_i]$. Usually, there are many such parameters, and thus, there are many possible combinations – even if we limit ourselves to extreme cases, when each parameter p_i is equal to either \underline{p}_i or to \bar{p}_i , we will still get 2^n possible combinations. It is therefore not feasible to physically check how the system behaves under all such combination. Instead, we need to rely on computer simulations.

Formulation of the problem. There are known techniques for using computer simulation to check that the system satisfies the given specifications for all possible combinations of these parameters. These techniques, however, have been originally designed for the case when we have an exact model of the system.

In principle, we can also use these techniques in more realistic situations, when the corresponding model is only approximate. However, as we show in this paper, the use of these techniques leads to overestimation of the corresponding uncertainty. We also show that a proper modification of these techniques leads to a drastic decrease of this overestimation and thus, to more accurate estimations.

HOW TO CHECK SPECIFICATIONS WHEN WE HAVE AN EXACT MODEL OF A SYSTEM: REMINDER

Case of an exact model: description. To run the corresponding computer simulations, we need to have a computer model that, given the values of the parameters p_1, \dots, p_n , estimates the corresponding value of the parameter q . Let us first consider situations when this computer model is exact, i.e., when this model enables us to compute the exact value q :

$$q = q(p_1, \dots, p_n). \quad (3)$$

In most engineering situations, deviations from nominal values are small. Usually, possible deviations $\Delta p_i \stackrel{\text{def}}{=} p_i - \tilde{p}_i$ from nominal values are reasonably small; see, e.g., [9]. In this paper, we will restrict ourselves to such situations.

In such situations, linearization is possible. In such situations, we can plug in the values $p_i = \tilde{p}_i + \Delta p_i$ into the formula (3), expand the resulting expression in Taylor series in terms of small values Δp_i , and ignore terms which are quadratic (or of higher order) in terms of Δp_i .

As a result, we get the following expression:

$$q(p_1, \dots, p_n) = q(\tilde{p}_1, \dots, \tilde{p}_n) + \sum_{i=1}^n \frac{\partial q}{\partial p_i} \cdot \Delta p_i, \quad (5)$$

or, equivalently,

$$q(p_1, \dots, p_n) = \tilde{q} + \sum_{i=1}^n c_i \cdot \Delta p_i, \quad (6)$$

where we denoted

$$\tilde{q} \stackrel{\text{def}}{=} q(\tilde{x}_1, \dots, \tilde{x}_n) \text{ and } c_i \stackrel{\text{def}}{=} \frac{\partial q}{\partial p_i}. \quad (7)$$

How to use the linearized model to check that specifications are satisfied: analysis of the problem. To make sure that we always have $q \leq t$, we need to guarantee that the largest possible value \bar{q} of the function q does not exceed t .

How can we compute this upper bound \bar{q} ? The maximum of the sum (6) is attained when each of n terms $c_i \cdot \Delta p_i$ attains the largest possible value. Each of these terms is a linear function of $\Delta p_i \in [-\Delta_i, \Delta_i]$. A linear function is always monotonic, and thus, it attains its largest value on an interval on one of its endpoint:

- When $c_i \geq 0$, the linear function $c_i \cdot \Delta p_i$ is increasing and thus, its largest value is attained when Δp_i is the largest, i.e., when $\Delta p_i = \Delta_i$. The resulting largest value of this linear function is $c_i \cdot \Delta_i$.
- When $c_i \leq 0$, the linear function $c_i \cdot \Delta p_i$ is decreasing and thus, its largest value is attained when Δp_i is the smallest, i.e., when $\Delta p_i = -\Delta_i$. The resulting largest value of this linear function is $c_i \cdot (-\Delta_i) = -c_i \cdot \Delta_i$.

In both cases, the largest value of the linear function is equal to $|c_i| \cdot \Delta_i$. Thus, the desired largest possible value \bar{q} of the quantity q is equal to

$$\bar{q} = \tilde{q} + \sum_{i=1}^n |c_i| \cdot \Delta_i; \quad (8)$$

see, e.g., [4, 9].

How to estimate the derivatives c_i ? Sometimes, we have an explicit formula for computing $q(p_1, \dots, p_n)$. In this case, by explicitly differentiating the corresponding expression, we can get formulas for computing the derivatives c_i .

In most real-life situations, however, there is no explicit formula. To find the value $q(p_1, \dots, p_n)$ corresponding to the parameter values p_1, \dots, p_n – e.g., to find the corresponding stress

– we need to solve a system of partial differential equations. In such situations, the dependence $q(p_1, \dots, p_n)$ is given in terms of a complex algorithm (and not an explicit formula), and thus, computing the derivative is not as straightforward.

Since we do not have an analytical expression for the derivative c_i , we need to use *numerical differentiation* to estimate c_i . The main idea behind numerical differentiation is to use the definition of the partial derivative

$$c_i = \frac{\partial q}{\partial p_i} \stackrel{\text{def}}{=} \lim_{h_i \rightarrow 0} \frac{q(\tilde{p}_1, \dots, \tilde{p}_{i-1}, \tilde{p}_i + h_i, \tilde{p}_{i+1}, \dots, \tilde{p}_n) - \tilde{q}}{h_i}. \quad (9)$$

The limit means that for small h_i , we have approximate equality, so we can estimate c_i as the ratio

$$c_i \approx \frac{q(\tilde{p}_1, \dots, \tilde{p}_{i-1}, \tilde{p}_i + h_i, \tilde{p}_{i+1}, \dots, \tilde{p}_n) - \tilde{q}}{h_i} \quad (10)$$

corresponding to some small value h_i .

What value h_i should we choose? We have assumed that when $|\Delta p_i| \leq \Delta_i$, then the dependence $q(p_1, \dots, p_n)$ can be safely linearized. Thus, when $|h_i| \leq \Delta_i$, the linearized formula (6) implies that

$$q(\tilde{p}_1, \dots, \tilde{p}_{i-1}, \tilde{p}_i + h_i, \tilde{p}_{i+1}, \dots, \tilde{p}_n) = \tilde{q} + h_i \cdot c_i, \quad (11)$$

and so, within this accuracy, the formula (10) is exact:

$$c_i = \frac{q(\tilde{p}_1, \dots, \tilde{p}_{i-1}, \tilde{p}_i + h_i, \tilde{p}_{i+1}, \dots, \tilde{p}_n) - \tilde{q}}{h_i}. \quad (12)$$

Substituting the formula (12) into the expression (8), we get

$$\bar{q} = \tilde{q} + \sum_{i=1}^n \frac{|q(\tilde{p}_1, \dots, \tilde{p}_{i-1}, \tilde{p}_i + h_i, \tilde{p}_{i+1}, \dots, \tilde{p}_n) - \tilde{q}|}{h_i} \cdot \Delta_i. \quad (13)$$

This formula is accurate for all the values h_i for which $|h_i| \leq \Delta_i$. It is therefore reasonable to select the value h_i that would decrease the number of computations.

No matter which values h_i we select, we need to run the simulated model $n + 1$ times:

- one time to compute the value $\tilde{q} = q(\tilde{p}_1, \dots, \tilde{p}_n)$, and then
- for each of the n parameters q_i , $1 \leq i \leq n$, to compute the value $q(\tilde{p}_1, \dots, \tilde{p}_{i-1}, \tilde{p}_i + h_i, \tilde{p}_{i+1}, \dots, \tilde{p}_n)$.

After that, we need n subtractions (between different values of q), n divisions (by h_i), n multiplications (by Δ_i) and n additions (to add up all the terms). Out of these arithmetic operations:

- addition and subtraction are the fastest,
- multiplication is somewhat longer – since multiplication contains several additions, and
- division is the longest, since it usually involves several multiplications.

Thus, to speed up computations, we need to select the values h_i that would allow us to avoid multiplication and division. One can easily see that this is possible when we take $h_i = \Delta_i$. In this case, the formula (13) takes a simplified form

$$\bar{q} = \tilde{q} + \sum_{i=1}^n |q_i - \tilde{q}|, \quad (14)$$

where we denoted

$$q_i \stackrel{\text{def}}{=} q(\tilde{p}_1, \dots, \tilde{p}_{i-1}, \tilde{p}_i + \Delta_i, \tilde{p}_{i+1}, \dots, \tilde{p}_n). \quad (15)$$

Thus, we arrive at the following technique (see, e.g., [4]).

How to use the linearized model to check that specifications are satisfied: resulting technique. We know:

- an algorithm $q(p_1, \dots, p_n)$ that, given the values of the parameters p_1, \dots, p_n , computes the value of the quantity q ;
- a threshold t that needs to be satisfied;
- for each parameter p_i , we know its nominal value \tilde{p}_i and the largest possible deviation Δ_i from this nominal value.

Based on this information, we need to check whether $q(p_1, \dots, p_n) \leq t$ for all possible combinations of values p_i from the corresponding intervals $[\tilde{p}_i - \Delta_i, \tilde{p}_i + \Delta_i]$.

We can perform this checking as follows:

- first, we apply the algorithm q to compute the value $\tilde{q} = q(\tilde{p}_1, \dots, \tilde{p}_n)$;
- then, for each i from 1 to n , we apply the algorithm q to compute the value $q_i = q(\tilde{p}_1, \dots, \tilde{p}_{i-1}, \tilde{p}_i + \Delta_i, \tilde{p}_{i+1}, \dots, \tilde{p}_n)$;
- after that, we compute $\bar{q} = \tilde{q} + \sum_{i=1}^n |q_i - \tilde{q}|$;
- finally, we check whether $\bar{q} \leq t$.

If $\bar{q} \leq t$, this means that the desired specifications are always satisfied. If $\bar{q} > t$, this means that for some combinations of possible values p_i , the specifications are not satisfied.

Possibility of a further speed-up. The formula (14) requires $n + 1$ calls to the program that computes q for given values of parameters. In many practical situations, the program q takes a reasonably long time to compute, and the number of parameters is large. In such situations, the corresponding computations require a very long time.

A possibility to speed up the corresponding computations comes from the properties of the Cauchy distribution, i.e., a distribution with a probability density function

$$\rho(x) = \frac{1}{\pi \cdot \Delta} \cdot \frac{1}{1 + \left(\frac{x}{\Delta}\right)^2}. \quad (16)$$

The possibility to use Cauchy distributions comes from the fact that they have the following property: if η_i are independent variables which are Cauchy distributed with parameters Δ_i , then for each tuple of real numbers c_1, \dots, c_n , the linear combination $\sum_{i=1}^n c_i \cdot \eta_i$ is also Cauchy distributed, with parameter $\Delta = \sum_{i=1}^n |c_i| \cdot \Delta_i$.

Thus, we can find Δ as follows [6]:

- first, for $k = 1, \dots, N$, we simulate random variables $\eta_i^{(k)}$ which are Cauchy-distributed with parameters Δ_i ;
- for each k , we then estimate $\Delta y^{(k)} = \sum_{i=1}^n c_i \cdot \eta_i^{(k)}$ as $\Delta y^{(k)} = y^{(k)} - \tilde{y}$, where

$$y^{(k)} = q(\tilde{p}_1 + \eta_1^{(k)}, \dots, \tilde{p}_n + \eta_n^{(k)}); \quad (17)$$

- based on the population of N values $\Delta y^{(1)}, \dots, \Delta y^{(N)}$ which is Cauchy-distributed with parameter Δ , we find this parameter;
- finally, we follow the formula (8) and compute $\bar{q} = \tilde{q} + \Delta$.

(see [6] for technical details).

In this Monte-Carlo-type technique, we need $N + 1$ calls to the program that computes q . The accuracy of the resulting estimate depends only on the sample size N and *not* on the number of inputs n . Thus, for a fixed desired accuracy, when n is sufficiently large, this method requires much fewer calls to q and is, thus, much faster. For example, if we want to estimate Δ with relative accuracy 20%, then we need $N = 100$ simulations, so for $n \gg 200$, this method is much faster than a straightforward application of the formula (14).

1 WHAT IF WE TAKE INTO ACCOUNT MODEL INACCURACY

Models are rarely exact. Engineering systems are usually complex. As a result, it is rarely possible to find explicit expressions for q as a function of the parameters p_1, \dots, p_n . Usually, we have some approximate computations. For example, if q is obtained by solving a system of partial differential equations, we use, e.g., the Finite Element method to find the approximate solution and thus, the approximate value of the quantity q .

How model inaccuracy is usually described. In most practical situations, at best, we know the upper bound ε on the accuracy of the computational model. In such cases, for each tuple of parameters p_1, \dots, p_n , once we apply the computational model and get the value $Q(p_1, \dots, p_n)$, the actual (unknown) value $q(p_1, \dots, p_n)$ of the quantity q satisfies the inequality

$$|Q(p_1, \dots, p_n) - q(p_1, \dots, p_n)| \leq \varepsilon. \quad (18)$$

How this model inaccuracy affects the above checking algorithms: analysis of the problem. Let us start with the formula (14). This formula assumes that we know the exact values of $\tilde{q} = q(\tilde{p}_1, \dots, \tilde{p}_n)$ and q_i (as defined by the formula (15)). Instead, we know the values

$$\tilde{Q} \stackrel{\text{def}}{=} Q(\tilde{p}_1, \dots, \tilde{p}_n) \quad (19)$$

and

$$Q_i \stackrel{\text{def}}{=} Q(\tilde{p}_1, \dots, \tilde{p}_{i-1}, \tilde{p}_i + \Delta_i, \tilde{p}_{i+1}, \dots, \tilde{p}_n) \quad (20)$$

which are ε -close to the values \tilde{q} and q_i . We can apply the formula (14) to these approximate values, and get

$$\bar{Q} = \tilde{Q} + \sum_{i=1}^n |Q_i - \tilde{Q}|. \quad (21)$$

Here, $|\tilde{Q} - \tilde{q}| \leq \varepsilon$ and $|Q_i - q_i| \leq \varepsilon$, hence $|(Q_i - \tilde{Q}) - (q_i - \tilde{q})| \leq 2\varepsilon$ and $||Q_i - \tilde{Q}| - |q_i - \tilde{q}|| \leq 2\varepsilon$. By adding up all these inequalities, we conclude that

$$|\bar{q} - \bar{Q}| \leq (2n + 1) \cdot \varepsilon. \quad (22)$$

Thus, the only information that we have about the desired upper bound \bar{q} is that $\bar{q} \leq B$, where

$$B \stackrel{\text{def}}{=} \bar{Q} + (2n + 1) \cdot \varepsilon. \quad (23)$$

Hence, we arrive at the following method.

How this model inaccuracy affects the above checking algorithms: resulting method. We know:

- an algorithm $Q(p_1, \dots, p_n)$ that, given the values of the parameters p_1, \dots, p_n , computes the value of the quantity q with a known accuracy ε ;

- a threshold t that needs to be satisfied;
- for each parameter p_i , we know its nominal value \tilde{p}_i and the largest possible deviation Δ_i from this nominal value.

Based on this information, we need to check whether $q(p_1, \dots, p_n) \leq t$ for all possible combinations of values p_i from the corresponding intervals $[\tilde{p}_i - \Delta_i, \tilde{p}_i + \Delta_i]$.

We can perform this checking as follows:

- first, we apply the algorithm Q to compute the value $\tilde{Q} = Q(\tilde{p}_1, \dots, \tilde{p}_n)$;
- then, for each i from 1 to n , we apply the algorithm Q to compute the value $Q_i = Q(\tilde{p}_1, \dots, \tilde{p}_{i-1}, \tilde{p}_i + \Delta_i, \tilde{p}_{i+1}, \dots, \tilde{p}_n)$;
- after that, we compute $B = \tilde{Q} + \sum_{i=1}^n |Q_i - \tilde{Q}| + (2n + 1) \cdot \varepsilon$;
- finally, we check whether $B \leq t$.

If $B \leq t$, this means that the desired specifications are always satisfied. If $B > t$, this means that we cannot guarantee that the specifications are always satisfied.

Comments.

- Please note that, in contrast to the case of the exact model, if $B > t$, this does not necessarily mean that the specifications are not satisfied: maybe they are satisfied, but we cannot check that since we only know approximate values of q .
- Similar bounds can be found for the estimates based on the Cauchy distribution.
- The above estimate B is not the best that we can get, but it has been proven that computing the best estimate would require un-realistic exponential time [3, 7] – i.e., time which grows as 2^s with the size s of the input; thus, when we only consider feasible algorithms, overestimation is inevitable.

Problem. When n is large, then, even for reasonably small inaccuracy ε , the value $(2n + 1) \cdot \varepsilon$ is large.

In this paper, we show how we can get better estimates for the difference between the desired bound \tilde{q} and the computed bound \tilde{Q} .

2 HOW TO GET BETTER ESTIMATES

Main idea. As we have mentioned earlier, usually, we know the partial differential equations that describe the engineering system. Model inaccuracy comes from the fact that we do not have an analytical solution to this system of equations, so we have to use numerical (approximate) methods.

Usual numerical methods for solving systems of partial differential equations involve discretization of space – e.g., the use of Finite Element Methods.

Strictly speaking, the resulting inaccuracy is deterministic. However, in most cases, for all practical purposes, this inaccuracy can be viewed as random:

- when we select a different combination of parameters,
- we get an unrelated value of discretization-based inaccuracy.

In other words, we can view the differences $Q(p_1, \dots, p_n) - q(p_1, \dots, p_n)$ corresponding to different tuples (p_1, \dots, p_n) as independent random variables. In particular, this means that the differences $\tilde{Q} - \tilde{q}$ and $Q_i - q_i$ are independent random variables.

Technical details. What is a probability distribution for these random variables?

All we know about each of these variables is that its values are located somewhere in the interval $[-\varepsilon, \varepsilon]$. We do not have any reason to assume that some values from this interval are more probable than others, so it is reasonable to assume that all the values are equally probable, i.e., that we have a uniform distribution on this interval.

For this uniform distribution, the mean is 0, and the standard deviation is $\sigma = \frac{\varepsilon}{\sqrt{3}}$.

Auxiliary idea: how to get a better estimate for \tilde{q} . In our main algorithm, we apply the computational model Q to $n + 1$ different tuples. What we suggest it to apply it to one more tuple (making it $n + 2$ tuples), namely, computing an approximation

$$M \stackrel{\text{def}}{=} Q(\tilde{p}_1 - \Delta_1, \dots, \tilde{p}_n - \Delta_n) \quad (24)$$

to the value

$$m \stackrel{\text{def}}{=} q(\tilde{p}_1 - \Delta_1, \dots, \tilde{p}_n - \Delta_n). \quad (25)$$

In the linearized case (6), one can easily check that

$$\tilde{q} + \sum_{i=1}^n q_i + m = (n + 2) \cdot \tilde{q}, \quad (26)$$

i.e.,

$$\tilde{q} = \frac{1}{n + 2} \cdot \left(\tilde{q} + \sum_{i=1}^n q_i + m \right). \quad (27)$$

Thus, we can use the following formula to come up with a new estimate \tilde{Q}_{new} for \tilde{q} :

$$\tilde{Q}_{\text{new}} = \frac{1}{n + 2} \cdot \left(\tilde{Q} + \sum_{i=1}^n Q_i + m \right). \quad (27a)$$

For the differences $\Delta\bar{q}_{\text{new}} \stackrel{\text{def}}{=} \bar{Q}_{\text{new}} - \bar{q}$, $\Delta\bar{q} \stackrel{\text{def}}{=} \bar{Q} - \bar{q}$, $\Delta\tilde{q} \stackrel{\text{def}}{=} \tilde{Q} - \tilde{q}$, $\Delta q_i \stackrel{\text{def}}{=} Q_i - q_i$, and $\Delta m \stackrel{\text{def}}{=} M - m$, we have the following formula:

$$\Delta\tilde{q}_{\text{new}} = \frac{1}{n+2} \cdot \left(\Delta\tilde{q} + \sum_{i=1}^n \Delta q_i + \Delta m \right). \quad (27b)$$

The left-hand side is the arithmetic average of $n+2$ independent identically distributed random variables, with mean 0 and variance $\sigma^2 = \frac{\varepsilon^2}{3}$. Hence (see, e.g., [10]), the mean of this average $\Delta\tilde{q}_{\text{new}}$ is the average of the means, i.e., 0, and the variance is equal to $\sigma^2 = \frac{\varepsilon^2}{3 \cdot (n+2)} \ll \frac{\varepsilon^2}{3} = \sigma^2[\Delta\tilde{q}]$.

Thus, this average \tilde{Q}_{new} is a more accurate estimation of the quantity \tilde{q} than \tilde{Q} .

Let us use this better estimate for \tilde{q} when estimating the upper bound \bar{q} . Since the average \tilde{Q}_{new} is a more accurate estimation of the quantity \tilde{q} than \tilde{Q} , let us use this average instead of \tilde{Q} when estimating \bar{Q} . In other words, instead of the estimate (21), let us use a new estimate

$$\bar{Q}_{\text{new}} = \tilde{Q}_{\text{new}} + \sum_{i=1}^n |Q_i - \tilde{Q}|. \quad (28)$$

Let us estimate the accuracy of this new approximation.

The formula (14) can be described in the following equivalent form:

$$\bar{q} = \tilde{q} + \sum_{i=1}^n s_i \cdot (q_i - \tilde{q}) = \left(1 - \sum_{i=1}^n s_i \right) \cdot \tilde{q} + \sum_{i=1}^n s_i \cdot q_i, \quad (29)$$

where $s_i \in \{-1, 1\}$ are the signs of the differences $q_i - \tilde{q}$.

Similarly, we get

$$\bar{Q}_{\text{new}} = \left(1 - \sum_{i=1}^n s_i \right) \cdot \tilde{Q}_{\text{new}} + \sum_{i=1}^n s_i \cdot Q_i. \quad (30)$$

Thus, for the difference $\Delta\bar{q} \stackrel{\text{def}}{=} \bar{Q}_{\text{new}} - \bar{q}$, we have

$$\Delta\bar{q}_{\text{new}} = \left(1 - \sum_{i=1}^n s_i \right) \cdot \Delta\tilde{q}_{\text{new}} + \sum_{i=1}^n s_i \cdot \Delta q_i. \quad (31)$$

Here, the differences $\Delta\tilde{q}_{\text{new}}$ and Δq_i are independent random variables. According to the Central Limit Theorem (see, e.g.,

[10]), for large n , the distribution of a linear combination of many independent random variables is close to Gaussian. The mean of the resulting distribution is the linear combination of the means, thus equal to 0.

The variance of a linear combination $\sum_i k_i \cdot \eta_i$ of independent random variables η_i with variances σ_i^2 is equal to $\sum_i k_i^2 \cdot \sigma_i^2$. Thus, in our case, the variance σ^2 of the difference $\Delta\bar{q}$ is equal to

$$\sigma^2 = \left(1 - \sum_{i=1}^n s_i \right)^2 \cdot \frac{\varepsilon^2}{3 \cdot (n+2)} + \sum_{i=1}^n \frac{\varepsilon^2}{3}. \quad (32)$$

Here, since $|s_i| \leq 1$, we have $\left| 1 - \sum_{i=1}^n s_i \right| \leq n+1$, so (32) implies that

$$\sigma^2 \leq \frac{\varepsilon^2}{3} \cdot \left(\frac{(n+1)^2}{n+2} + n \right). \quad (33)$$

Here, $\frac{(n+1)^2}{n+2} \leq \frac{(n+1)^2}{n+1} = n+1$, hence

$$\sigma^2 \leq \frac{\varepsilon^2}{3} \cdot (2n+1). \quad (33)$$

For a normal distribution, with almost complete certainty, all the values are concentrated within k_0 standard deviations away from the mean: within 2σ with confidence 0.95, within 3σ with degree of confidence 0.999, within 6σ with degree of confidence $1 - 10^{-8}$. Thus, we can safely conclude that

$$\bar{q} \leq \bar{Q}_{\text{new}} + k_0 \cdot \sigma \leq \bar{Q}_{\text{new}} + k_0 \cdot \frac{\varepsilon}{\sqrt{3}} \cdot \sqrt{2n+1}. \quad (34)$$

Here, inaccuracy grows as $\sqrt{2n+1}$, which is much better than in the traditional approach, where it grows proportionally to $2n+1$ – and we achieve this drastic reduction of the overestimation, basically by using one more run of the program Q in addition to the previously used $n+1$ runs.

So, we arrive at the following method.

Resulting method. We know:

- an algorithm $Q(p_1, \dots, p_n)$ that, given the values of the parameters p_1, \dots, p_n , computes the value of the quantity q with a known accuracy ε ;
- a threshold t that needs to be satisfied;
- for each parameter p_i , we know its nominal value \tilde{p}_i and the largest possible deviation Δ_i from this nominal value.

Based on this information, we need to check whether $q(p_1, \dots, p_n) \leq t$ for all possible combinations of values p_i from the corresponding intervals $[\tilde{p}_i - \Delta_i, \tilde{p}_i + \Delta_i]$.

We can perform this checking as follows:

- first, we apply the algorithm Q to compute the value $\tilde{Q} = Q(\tilde{p}_1, \dots, \tilde{p}_n)$;
- then, for each i from 1 to n , we apply the algorithm Q to compute the value $Q_i = Q(\tilde{p}_1, \dots, \tilde{p}_{i-1}, \tilde{p}_i + \Delta_i, \tilde{p}_{i+1}, \dots, \tilde{p}_n)$;
- then, we compute $M = Q(\tilde{p}_1 - \Delta_1, \dots, \tilde{p}_n - \Delta_n)$;
- compute $\tilde{Q}_{\text{new}} = \frac{1}{n+2} \cdot \left(\tilde{Q} + \sum_{i=1}^n Q_i + M \right)$;
- compute $b = \tilde{Q}_{\text{new}} + \sum_{i=1}^n |Q_i - \tilde{Q}_{\text{new}}| + k_0 \cdot \sqrt{2n+1} \cdot \frac{\varepsilon}{\sqrt{3}}$, where k_0 depends on the level of confidence that we can achieve;
- finally, we check whether $b \leq t$.

If $b \leq t$, this means that the desired specifications are always satisfied. If $b > t$, this means that we cannot guarantee that the specifications are always satisfied.

Comment. For the Cauchy method, similarly, after computing $\tilde{Q} = Q(\tilde{p}_1, \dots, \tilde{p}_n)$ and $Y^{(k)} = Q(\tilde{p}_1 + \eta_1^{(k)}, \dots, \tilde{p}_n + \eta_n^{(k)})$, we can compute the improved estimate \tilde{Q}_{new} for \tilde{q} as

$$\tilde{Q}_{\text{new}} = \frac{1}{N+1} \cdot \left(\tilde{Q} + \sum_{k=1}^N Y^{(k)} \right), \quad (35)$$

and estimate the parameter Δ based on the more accurate differences $\Delta y_{\text{new}}^{(k)} = Y^{(k)} - \tilde{Q}_{\text{new}}$.

Experimental testing. We tested our approach on the example of the seismic inverse problem in geophysics, where we need to reconstruct the velocity of sound at different spatial locations and at different depths based on the times that it takes for a seismic signal to get from the set-up explosion to different seismic stations. In this reconstruction, we used (a somewhat improved version of) the finite element technique that was originated by John Hole [2]; the resulting techniques are described in [1, 5, 8].

In [1, 5, 8], we used the formula (14) and the Cauchy-based techniques to estimate how the measurement uncertainty affects the results of data processing. To test our method, we used the above formulas to compute the improved values \tilde{Q}_{new} . These improved values indeed lead to a better fit with data than the original values \tilde{Q} .

FUTURE WORK: CAN WE FURTHER IMPROVE THE ACCURACY?

How to improve the accuracy: a straightforward approach. As we have mentioned, the inaccuracy $Q \neq q$ is caused by the fact

that we are using a Finite Element method with a finite size elements. This inaccuracy comes from the fact that we ignore the difference between the values of the corresponding parameters within each element. For elements of linear size h , this inaccuracy Δx is proportional to $x' \cdot h$, where x' is the spatial derivative of x . In other words, the inaccuracy is proportional to the linear size h .

A straightforward way to improve the accuracy is to decrease h . For example, if we reduce h to $\frac{h}{2}$, then we decrease the resulting model inaccuracy ε to $\frac{\varepsilon}{2}$.

This decrease requires more computations. The number of computations is, crudely speaking, proportional to the number of elements. Since the elements fill the original area, and each element has volume h^3 , the number of such elements is proportional to h^{-3} .

So, if we go from the original value h to the smaller value h' , then we increase the number of computations by a factor of $K \stackrel{\text{def}}{=} \frac{h^3}{(h')^3}$.

This leads to decreasing the inaccuracy by a factor of $\frac{h}{h'}$, which is equal to $\sqrt[3]{K}$.

For example, in this straightforward approach, if we want to decrease the accuracy in half ($\sqrt[3]{K} = 2$), we will have to increase the number of computation steps by a factor of $K = 8$.

An alternative approach: description. An alternative approach is as follows. We select K small vectors $(\Delta_1^{(k)}, \dots, \Delta_n^{(k)})$, $1 \leq k \leq K$, which add up to 0. For example, we can arbitrarily select the first $K-1$ vectors and take $\Delta p_i^{(K)} = -\sum_{k=1}^{K-1} \Delta_i^{(k)}$.

Then, every time we need to estimate the value $q(p_1, \dots, p_n)$, instead of computing $Q(p_1, \dots, p_n)$, we compute the average

$$Q_K(p_1, \dots, p_n) = \frac{1}{K} \cdot \sum_{k=1}^K Q\left(p_1 + \Delta_1^{(k)}, \dots, p_n + \Delta_n^{(k)}\right). \quad (36)$$

Why this approach decreases inaccuracy. We know that $Q(p_1 + \Delta p_1, \dots, p_n + \Delta p_n) = q(p_1 + \Delta p_1, \dots, p_n + \Delta p_n) + \delta q$, where, in the small vicinity of the original tuple (p_1, \dots, p_n) :

- the expression $q(p_1 + \Delta p_1, \dots, p_n + \Delta p_n)$ is linear, and
- the differences δq are independent random variables with 0 mean.

Thus, we have

$$Q_K(p_1, \dots, p_n) = \frac{1}{K} \cdot \sum_{k=1}^K q\left(p_1 + \Delta_1^{(k)}, \dots, p_n + \Delta_n^{(k)}\right) +$$

$$\frac{1}{K} \cdot \sum_{k=1}^K \Delta q^{(k)}. \quad (37)$$

Due to linearity and the fact that $\sum_{k=1}^K \Delta_i^{(k)} = 0$, the first average in (37) is equal to $q(p_1, \dots, p_n)$. The second average is the average of K independent identically distributed random variables, and we have already recalled that this averaging decreases the inaccuracy by a factor of \sqrt{K} .

Thus, in this alternative approach:

- we increase the amount of computations by a factor of K , and
- as a result, we decrease the inaccuracy by a factor of \sqrt{K} .

The new approach is better than the straightforward one. In general, $\sqrt{K} > \sqrt[3]{K}$. Thus, with the same increase in computation time, the new method provides a better improvement in accuracy than the straightforward approach.

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REFERENCES

- [1] M. G. Averill, *Lithospheric Investigation of the Southern Rio Grande Rift*, University of Texas at El Paso, Department of Geological Sciences, PhD Dissertation, 2007.
- [2] J. A. Hole, “Nonlinear high-resolution three-dimensional seismic travel time tomography, *Journal of Geophysical Research*, 192, Vol. 97, No. B5, pp. 6553–6562.
- [3] V. Kreinovich, “Error estimation for indirect measurements is exponentially hard”, *Neural, Parallel, and Scientific Computations*, 1994, Vol. 2, No. 2, pp. 225–234.
- [4] V. Kreinovich, “Interval Computations and Interval-Related Statistical Techniques: Tools for Estimating Uncertainty of the Results of Data Processing and Indirect Measurements”, In: F. Pavese and A. B. Forbes (eds.), *Data Modeling for Metrology and Testing in Measurement Science*, Birkhauser-Springer, Boston, 2009, pp. 117–145.
- [5] V. Kreinovich, J. Beck, C. Ferregut, A. Sanchez, G. R. Keller, M. Averill, and S. A. Starks, “Monte-Carlo-type techniques for processing interval uncertainty, and their potential engineering applications”, *Reliable Computing*, 2007, Vol. 13, No. 1, pp. 25–69.
- [6] V. Kreinovich and S. Ferson, “A New Cauchy-Based Black-Box Technique for Uncertainty in Risk Analysis”, *Reliability Engineering and Systems Safety*, 2004, Vol. 85, No. 1–3, pp. 267–279.
- [7] V. Kreinovich, A. Lakeyev, J. Rohn, and P. Kahl, *Computational Complexity and Feasibility of Data processing and Interval Computations*, Kluwer, Dordrecht, 1997.
- [8] P. Pinheiro da Silva, A. Velasco, M. Ceberio, C. Servin, M. G. Averill, N. Del Rio, L. Longpré, and V. Kreinovich, “Propagation and provenance of probabilistic and interval uncertainty in cyberinfrastructure-related data processing and data fusion”, In: R. L. Muhanna and R. L. Mullen (eds.), *Proceedings of the International Workshop on Reliable Engineering Computing REC’08*, Savannah, Georgia, February 20–22, 2008, pp. 199–234.
- [9] S. Rabinovich, *Measurement Errors and Uncertainties: Theory and Practice*, Springer Verlag, New York, 2005.
- [10] D. J. Sheskin, *Handbook of Parametric and Nonparametric Statistical Procedures*, Chapman & Hall/CRC, Boca Raton, Florida, 2011.